

# Association between *SERPING1* rs2511989 polymorphism and age-related macular degeneration: Meta-analysis

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## Abstract

• **AIM:** To investigate the association between *SERPING1* rs2511989 (G>A) polymorphism and age-related macular degeneration (AMD).

• **METHODS:** A number of electronic databases (up to July 15, 2014) were searched independently by two investigators. A Meta-analysis was performed on the association between *SERPING1* rs2511989 polymorphism and AMD. Pooled odds ratios (ORs) with 95% confidence intervals (CIs) were estimated.

• **RESULTS:** Eight studies with 16 cohorts consisting of 9163 cases and 6813 controls were included in this Meta-analysis. There was no significant association between rs2511989 polymorphism and AMD under all genetic models in overall estimates (A vs G: OR = 0.938, 95% CI = 0.858–1.025; AA vs GG: OR = 0.871, 95% CI = 0.719–1.056; AG vs GG: OR = 0.944, 95% CI = 0.845–1.054; AA + AG vs GG: OR = 0.927, 95% CI = 0.823–1.044; AA vs AG + GG: OR = 0.890, 95% CI = 0.780–1.034). Cumulative Meta-analyses also showed a trend of no association between rs2511989 polymorphism and AMD as information accumulated by year. Subgroup analysis and Meta-regression analysis indicated that age-matching status was the main source of heterogeneity. Sensitivity analysis found the results in overall comparisons and subgroup comparisons of white subjects under the allele model were found to have significantly statistical differences after studies deviating from Hardy-Weinberg equilibrium (HWE) were excluded (overall: OR = 0.918, 95% CI = 0.844–0.999,  $P = 0.049$ ; whites: OR = 0.901, 95% CI = 0.817–0.994,  $P = 0.038$ ). However, the results were not

sufficiently robust for further sensitivity analysis and statistical differences disappeared on applying Bonferroni correction (with a significance level set at 0.05/25).

• **CONCLUSION:** This Meta-analysis indicates that *SERPING1* rs2511989 polymorphism and AMD tend to have no association with each other. Age matching status is a big confounding factor, and more studies with subtle designs are warranted in future.

• **KEYWORDS:** age-related macular degeneration; *SERPING1*; single nucleotide polymorphism; Meta-analysis

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## INTRODUCTION

Age-related macular degeneration (AMD) is a leading cause of irreversible blindness in older populations in developed countries, and has become a major public health issue<sup>[1,2]</sup>. The prevalence of AMD increases strongly with age, affecting 4% of the population over the age of 60, 10% of individuals older than 75, and 64% of the population after the age of 80<sup>[2-4]</sup>. In late-stage AMD, the disease is characterized by drusen (focal depositions of extracellular material in the Bruch's membrane beneath the retinal pigment epithelium) and usually without clinical symptoms<sup>[5,6]</sup>. In late-stage AMD, vision-threatening complications of choroidal neovascularization or atrophy develop, which can lead to severe irreversible central vision loss<sup>[7]</sup>.

Although many previous studies have identified that AMD is a complex disorder caused by the interaction of multiple genetic and environmental risk factors<sup>[8-12]</sup>, the specific cause still remains to be further investigated, to improve preclinical prediction. So far, remarkable progress has recently been made in understanding the genetic risk factors AMD. Several complement component and regulator genes have been identified, and these highlight the importance of the complement pathway in the pathogenesis of AMD. At least five complement genes have been found to be associated with the development of this disorder, including complement regulator factor H complement factor I, the complement components C3, the complement components C2, and factor B<sup>[13-16]</sup>.

**Table 1** Principle characteristics of the studies included in the Meta-analysis

Frist author	Cohorts	Year	Ethnicity	Genotyping	Case					Control					Control HWE(P)
					Size	GG	GA	AA	MAF	Size	GG	GA	AA	MAF	
Ennis <i>et al</i> <sup>[18]</sup>	UK	2008	Caucasians	Illumina	479	191	215	70	0.37	479	132	236	109	0.48	0.858
Ennis <i>et al</i> <sup>[18]</sup>	U (Lowa)	2008	Caucasians	Taqman	248	100	122	26	0.35	252	79	124	49	0.44	0.978
Park <i>et al</i> <sup>[19]</sup>	Mayo	2009	Caucasians	Taqman	476	179	211	80	0.39	310	103	157	50	0.41	0.445
Park <i>et al</i> <sup>[19]</sup>	Areds	2009	Caucasians	Illumina	1221	436	563	222	0.41	295	115	127	53	0.39	0.088
Allikmets <i>et al</i> <sup>[20]</sup>	Columbia	2009	Caucasians	Taqman	1004	449	431	124	0.34	363	151	171	41	0.35	0.476
Allikmets <i>et al</i> <sup>[20]</sup>	Lowa	2009	Caucasians	Combination	368	116	178	74	0.44	115	37	59	19	0.42	0.578
Allikmets <i>et al</i> <sup>[20]</sup>	Amsterdam	2009	Caucasians	Taqman	338	107	184	47	0.41	257	84	131	42	0.42	0.447
Allikmets <i>et al</i> <sup>[20]</sup>	Rotterdam	2009	Caucasians	Illumina	1017	328	518	171	0.42	842	285	407	150	0.42	0.822
Allikmets <i>et al</i> <sup>[20]</sup>	Australia	2009	Caucasians	Combination	741	251	367	123	0.41	327	105	157	65	0.44	0.649
Allikmets <i>et al</i> <sup>[20]</sup>	Germany	2009	Caucasians	Taqman	998	377	485	136	0.38	725	284	341	100	0.37	0.883
Allikmets <i>et al</i> <sup>[20]</sup>	Areds	2009	Caucasians	Taqman	415	133	188	94	0.45	213	90	81	42	0.39	0.004
Carter and Churchill <sup>[21]</sup>	UK	2011	Caucasians	direct sequencing	94	38	39	17	0.39	95	29	52	14	0.42	0.232
Lee <i>et al</i> <sup>[22]</sup>	US	2010	Caucasians	MassARRAY	556	213	273	70	0.37	256	74	135	47	0.45	0.287
Lu <i>et al</i> <sup>[23]</sup>	China	2010	Asians	SNaPshot	272	198	57	5	0.13	285	215	63	3	0.12	0.494
Nakata <i>et al</i> <sup>[24]</sup>	Japan	2011	Asians	Taqman+Illumina	401	293	102	6	0.14	1530	1107	384	37	0.15	0.591
Tian <i>et al</i> <sup>[25]</sup>	China	2012	Asians	MassARRAY	535	422	96	13	0.11	469	371	86	7	0.11	0.436

Combination: SSCP+direct sequencing+Affymetrix; MassARRAY: Sequenom MassARRAY technology; NA: Data not available; HWE: Hardy-Weinberg equilibrium.

The *SERPING1* gene encodes complement 1 inhibitor (C1INH), which is a glycoprotein that inhibits complement activation by interfering with the proteolytic activity of C1r/C1s in the classical pathway and mannose-binding protein-associated serine proteases in the lectin pathway<sup>[17]</sup>. As a member of the classical and lectin complement pathway, *SERPING1* is a plausible candidate gene for AMD. Recently, Ennis *et al*<sup>[18]</sup> reported a protective effect on AMD for the minor allele of a single nucleotide polymorphism (SNP rs2511989 G>A) within intron 6 of the *SERPING1* gene encoding for C1INH. Since then, many replication studies have been successively completed.

Although reported studies have focused on the association between the SNP rs2511989 and AMD, the results are contradictory and inconclusive. Hence, we performed a Meta-analysis based on eight candidate eligible studies consisting of 9163 cases and 6813 controls, which may confirm the association between the SNP rs2511989 in *SERPING1* gene and AMD. Additional SNP variants in the *SERPING1* gene that have been identified relating to AMD, such as rs2244169, rs2511990, rs2509897, rs1005510, and rs2511988, are not discussed in this article, because of a lack of studies.

**MATERIALS AND METHODS**

**Search Strategy** To identify the studies eligible for systematic review and Meta-analysis, the following electronic databases were searched: PubMed, Embase, Web of Science, Wanfang (Chinese), and the China National Knowledge Infrastructure database (CNKI), up to July 15, 2014. The search algorithm is as follows: (*SERPING1* OR complement component 1 inhibitor gene) AND (gene OR polymorphism\* OR variant\*OR single nucleotide polymorphism OR SNP) AND (age related macular degeneration OR macular degeneration OR AMD) AND (Case-control Studies OR

Case-Base Studies OR Case-Comparison Studies OR Case-Referent Studies). Additional studies were identified through a manual search of the references of the original studies and review articles. We retrieved through electronic searches to identify studies not yet included in the computerized databases. No language restriction in the search process and all studies performed on human subjects were included in our search.

**Inclusion/Exclusion Criteria** Studies included in our Meta-analysis had to meet the following inclusion criteria: 1) the original major objective was to explore the relationship between *SERPING1* gene polymorphism and AMD; 2) the studies were designed on the basis of independent case-control study with available data of genotype distributions and sufficient data for estimating odds ratios (ORs) with 95% confidence intervals (CIs); 3) as for the duplicated articles, the latest or the largest one was adopted. If several different cohorts were reported in the same article, they were considered as independent studies.

**Data Extraction** Data were collected carefully and independently by two independent investigators (Fang XY and Dong Y). The characteristics of the selected articles were shown in Table 1, including first author's name, cohorts, publication year, ethnicity, genotyping method, number of cases and controls, distribution of genotype frequency, minor allele frequency (MAF), and Hardy-Weinberg equilibrium (HWE) of control group (*P* value). Besides, we also listed the demographic characteristics of the study population and the age-matching status. When detailed data of mean ages were not available, we identified whether the age was matched or not according to claims on age matching status in the studies. If there was no claim, we treated cohorts for which the difference of the mean age was greater than 5y as age-unmatched, and vice versa<sup>[18,19]</sup>. Details are shown in

**Table 2 Demographic characteristics of study population**

Frist author	Cohorts	Age (mean±SD or mean)		Age matching status	Sex (male %)	
		Case	Control		Case	Control
Ennis <i>et al</i> <sup>[18]</sup>	UK	77.85±8.83	70.59±9.35	unmatched	37.9	48.5
Ennis <i>et al</i> <sup>[18]</sup>	US (Lowa)	81.18±9.12	74±9.04	unmatched	39.4	46.5
Park <i>et al</i> <sup>[19]</sup>	Mayo	76.9±9.6	69.5±8.2	unmatched	35.5	45.4
Park <i>et al</i> <sup>[19]</sup>	Areds	79.9±5.1	77.6±4.3	matched	40.5	44.1
Allikmets <i>et al</i> <sup>[20]</sup>	Columbia	NA	NA	matched	NA	NA
Allikmets <i>et al</i> <sup>[20]</sup>	Lowa	NA	NA	matched	NA	NA
Allikmets <i>et al</i> <sup>[20]</sup>	Amsterdam	NA	NA	matched	NA	NA
Allikmets <i>et al</i> <sup>[20]</sup>	Rotterdam	NA	NA	matched	NA	NA
Allikmets <i>et al</i> <sup>[20]</sup>	Australia	NA	NA	matched	NA	NA
Allikmets <i>et al</i> <sup>[20]</sup>	Germany	NA	NA	matched	NA	NA
Allikmets <i>et al</i> <sup>[20]</sup>	Areds	NA	NA	matched	NA	NA
Carter and Churchill <sup>[21]</sup>	UK	NA	NA	matched	28	33.7
Lu <i>et al</i> <sup>[23]</sup>	China	68.2±9.8	68.4±7.2	matched	46.3	46.3
Lee <i>et al</i> <sup>[22]</sup>	US	79.3	69.5	unmatched	31.9	45.3
Nakata <i>et al</i> <sup>[24]</sup>	Japan	77.38±8.39	NA	unmatched	71.6	41.5
Tian <i>et al</i> <sup>[25]</sup>	China	NA	NA	unmatched	60.6	46.3

NA: Data not available; SD: Standard deviation.

Table 2. In our study, ethnicities were subgrouped as Asians and whites. Disagreements were settled by discussion or consensus involving a third reviewer (Shi XF) when required.

**Statistical Analysis** Pooled ORs with corresponding 95% CIs were calculated to evaluate the strength of relationship between the *SERPING1* gene and AMD for the following five genetic models: the allele model (A vs G), the homozygote model (AA vs GG), the heterozygote model (AG vs GG), the dominant model (AA+AG vs GG), and the recessive model (AA vs AG+GG). A  $Z$  test was used to assess the significance of the pooled OR, in which  $P < 0.05$  was considered statistically significant. The Q test and  $I^2$  statistics were employed to evaluate between-study heterogeneity. If  $P_Q \leq 0.10$  or  $I^2 > 50\%$ , which indicated significant heterogeneity in the comparison models among studies<sup>[26]</sup>, the estimated pooled ORs for each study were calculated using a random-effects model (DerSimonian and Laird method)<sup>[27]</sup>. Otherwise, the fixed-effects model was considered more suitable (Mantel-Haenszel method)<sup>[28]</sup>. We also performed a cumulative Meta-analysis to provide a framework for updating a genetic effect from all studies, to measure how much of the genetic effect changed as evidence accumulated, and to find the trend in estimated risk effect<sup>[29]</sup>. For the cumulative Meta-analysis, studies were sorted chronologically by publication year.

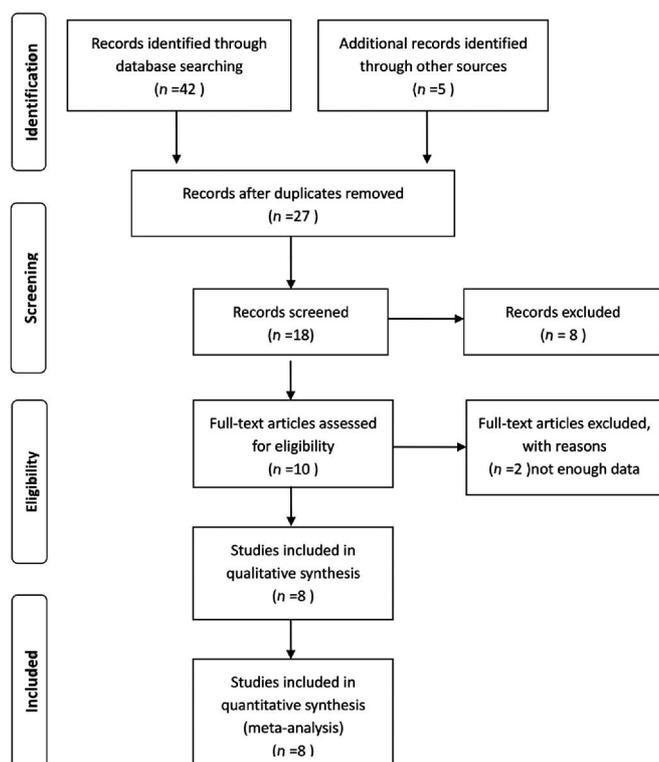
We performed subgroup and meta-regression analyses to explore potential sources of heterogeneity. All studies were classified based on ethnicity and age-matching status into white and Asian subgroups and age-matched and age-unmatched subgroups, respectively. We utilized univariate and multivariate meta-regression models to conduct the meta-regression. To tackle the issue of multiple

testing, 10 000 permutations of Monte Carlo simulation were necessary to adjust the result of the multivariate meta-regression model. The reliability of the results was assessed by sensitivity analysis performed by sequentially omitting individual studies or by excluding studies deviating from HWE. HWE was checked using a Chi-square test in each control group of the included studies, with results of  $P < 0.05$  considered as significantly deviating from HWE. Furthermore, to reduce the false positive error rate, the Bonferroni method was used to adjust the results of multiple comparisons. Because 25 comparisons were made in this Meta-analysis, the  $P$  value, which was less than  $0.05/25$  (0.002), indicated statistical significance after Bonferroni correction.

## RESULTS

**Study Characteristics** Figure 1 shows the selection process of this study. The initial search strategy identified 47 studies. Of these, 37 were excluded (20 were duplicate studies, 9 were on unrelated topics, and 8 were not case-control studies), leaving 10 studies for full review. Of these, two articles lacked sufficient data to estimate the OR and the 95% CI. Ultimately, eight studies met the inclusion criteria and were included in this Meta-analysis. We treated several different cohorts, which were reported in the same articles as independent studies. Therefore, there were 16 cohort studies consisting of 9163 cases and 6813 controls in our study. The detailed characteristics of the included studies are listed in Tables 1 and 2.

The genotype distributions in all controls were consistent with HWE, except the Age-related Eye Disease Study (AREDS) cohorts of Allikmets' study<sup>[20]</sup>. Among the 16 cohorts included, there were 3 studies of Asians<sup>[23-25]</sup> and 13



**Figure 1** PRISMA flow diagram of studies included in the Meta-analysis.

studies of white subjects<sup>[18-22]</sup>. In terms of age-matching status, there were six age-unmatched studies and ten age-matched studies. The rs2511989 SNP was detected by different assays, as shown in Table 1. The MAF for the rs2511989 SNP varied substantially between studies, from 0.35 to 0.45 in both AMD patients (average 0.40) and control subjects (average 0.42) among the white subgroup. In addition, in the Asian subgroup, the MAF ranged from 0.11 to 0.14 in AMD patients (average 0.13) and from 0.11 to 0.15 in control subjects (average 0.13), respectively.

**Overall Comparisons and Cumulative Meta-analysis** As a result of significant between-study heterogeneity detected in all genetic models for overall analysis (Table 3, Figure 2), a random-effects model was utilized to calculate pooled estimates. Overall, no significant relationship between the SNP rs2511989 and AMD was found in any genetic model (A vs G: OR=0.938, 95%CI=0.858-1.025; AA vs GG: OR=0.871, 95%CI=0.719-1.056; AG vs GG: OR=0.944, 95%CI=0.845-1.054; AA+AG vs GG: OR=0.927, 95%CI=0.823-1.044; AA vs AG+GG: OR=0.890, 95%CI=0.780-1.034). The cumulative Meta-analysis showed a trend of no association between SNP rs2511989 and AMD as information accumulated by year (Figure 3).

**Heterogeneity Analysis** Subgroup analyses were conducted to investigate potential sources of heterogeneity. Among the four subgroups, there was no significant association apart from the age-unmatched subgroup, which showed significant protective associations between SNP rs2511989 and AMD in all genetic models (A vs G: OR=0.813, 95%CI=0.695-0.951;

AA vs GG: OR=0.619, 95%CI=0.434-0.883; AG vs GG: OR=0.812, 95%CI=0.715-0.921; AA+AG vs GG: OR=0.771, 95%CI=0.634-0.939; AA vs AG+GG: OR=0.716, 95%CI=0.528-0.973). We found that the heterogeneity in the Asian subgroup and the age-matched subgroup either disappeared or was moderate. However, there was still significant heterogeneity in the white subgroup and the age-unmatched subgroup. Therefore, ethnicity and age-matching status might be a main source of heterogeneity. Table 3 and Figure 4 showed detailed results.

Univariate and multivariate meta-regression were employed to further explore the influence of ethnicity, age-matching status and genotyping on heterogeneity (Table 4). These results indicated that the majority of the heterogeneity comes from the age-matching status (adjusted  $P=0.005<0.05$ ), and not ethnicity or genotyping. Moreover, age-matching status could explain 67.0% of the heterogeneity, according to univariate meta-regression.

**Sensitivity Analysis** Sensitivity analysis was conducted to evaluate the stability of the results. Using a stepwise process, Meta-analysis was performed repeatedly, with each particular study sequentially excluded. The results of sensitivity analysis suggested that no single study could influence ORs in overall comparisons and subgroup analysis, except the AREDS cohorts of Allikmets' study (Figure 5). Coincidentally, the AREDS cohort was the only study that deviated from HWE in controls. After excluding this cohort, significant associations were observed for overall comparisons and white subgroup comparisons under the allele model (overall: OR=0.918, 95%CI=0.844-0.999,  $P=0.049$ ; whites: OR=0.901, 95%CI=0.817-0.994,  $P=0.038$ ), and the results of the age-matching status subgroup analysis were not changed. However, we found that these associations were not robust after we performed a sensitivity analysis on the remaining 15 studies (Figure 5F). Some of the results were materially altered and suggested different conclusion, as each individual study was sequentially omitted. Moreover, when we applied Bonferroni correction (significance set at 0.05/20) the association did not survive.

**Publication Bias** The Begg's funnel plot and Egger's test were performed for the overall comparisons and subgroup analysis. No obvious visual asymmetry was observed in Begg's funnel plots, and all the  $P$  values of the Egger's test were greater than 0.05, indicating no statistical evidence for publication bias among studies (Table 3, Figure 6).

## DISCUSSION

As is well-known, the hypothesis that the complement system participates in the pathology of AMD has been supported by many researchers in a number of disciplines, including animal models of neovascularization, histopathology and genetics<sup>[30]</sup>. As an inhibitor of the classical and lectin pathways of complement activation, *SERPINC1* expression could plausibly be protective against complement injury in

**Table 3 Summary risk estimates for association between SNP rs2511989 and AMD**

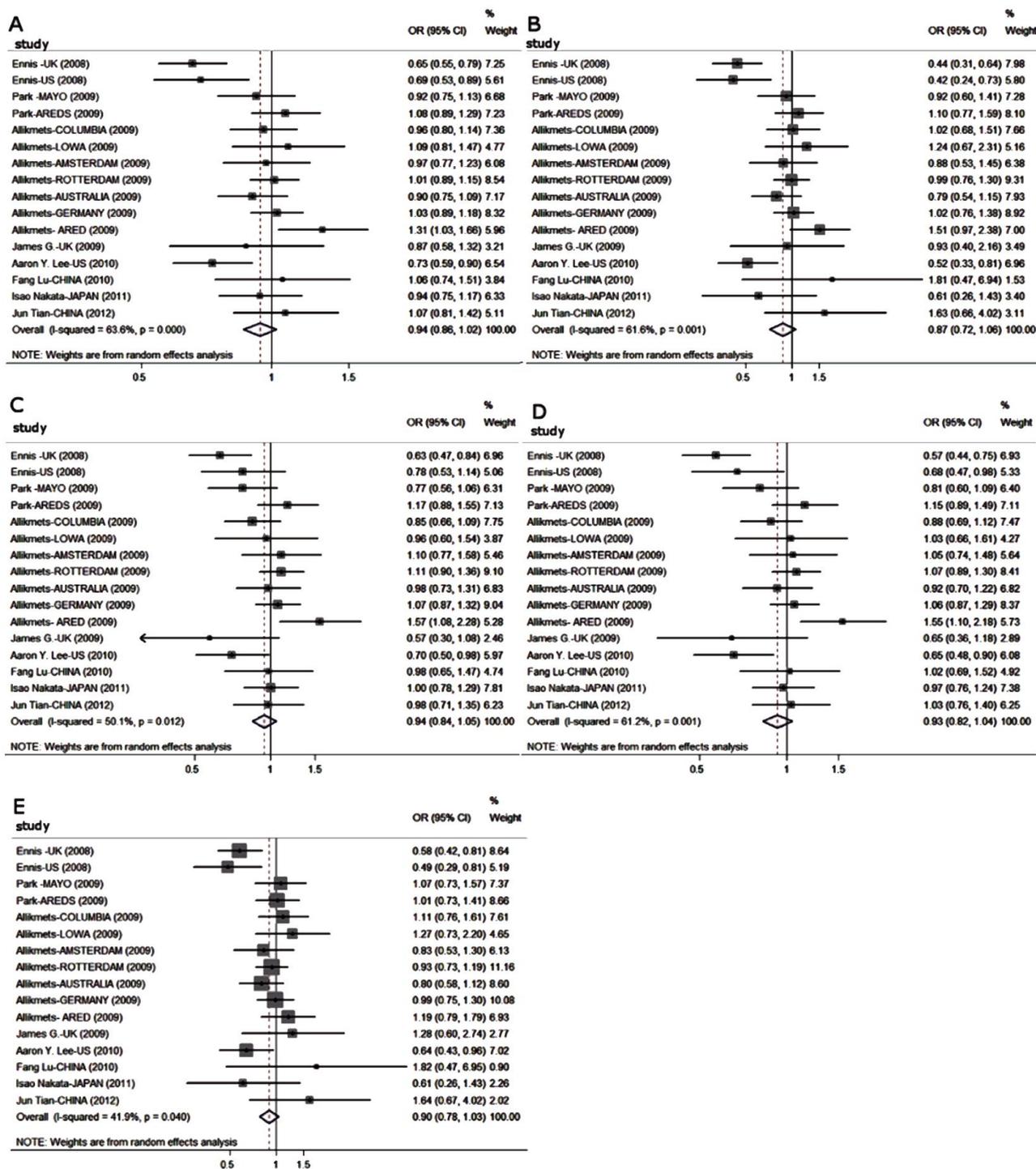
Stratifications	Studies (n)	Model	Pooled estimate		Heterogeneity		Egger's test P
			OR (95%CI)	P <sub>Z</sub>	I <sup>2</sup> (%)	P <sub>Q</sub>	
<b>Overall</b>							
A vs G	16	R	0.938 (0.858-1.025)	0.155	63.6	<0.001	0.854
AA vs GG	16	R	0.871 (0.719-1.056)	0.159	61.6	0.001	0.948
AG vs GG	16	R	0.944 (0.845-1.054)	0.306	50.1	0.012	0.280
AA+AG vs GG	16	R	0.927 (0.823-1.044)	0.211	61.2	0.001	0.387
AA vs AG+GG	16	R	0.890 (0.780-1.034)	0.135	41.9	0.040	0.595
<b>Asians</b>							
A vs G	3	F	1.002 (0.857-1.171)	0.984	0.0	0.725	0.615
AA vs GG	3	F	1.042 (0.604-1.800)	0.882	30.6	0.237	0.407
AG vs GG	3	F	0.993 (0.830-1.187)	0.935	0.0	0.993	0.352
AA+AG vs GG	3	F	0.998 (0.839-1.187)	0.980	0.0	0.948	0.435
AA vs AG+GG	3	F	1.043 (0.605-1.797)	0.880	31.6	0.232	0.614
<b>Caucasians</b>							
A vs G	13	R	0.925 (0.835-1.024)	0.134	70.0	<0.001	0.603
AA vs GG	13	R	0.853 (0.697-1.045)	0.124	66.4	<0.001	0.621
AG vs GG	13	R	0.930 (0.811-1.067)	0.300	59.8	0.003	0.310
AA+AG vs GG	13	R	0.908 (0.785-1.051)	0.195	68.4	<0.001	0.360
AA vs AG+GG	13	R	0.889 (0.769-1.028)	0.112	46.7	0.032	0.956
<b>Age-matched</b>							
A vs G	10	F	1.018 (0.957-1.083)	0.571	0.0	0.562	0.802
AA vs GG	10	F	1.029 (0.904-1.172)	0.662	0.0	0.713	0.388
AG vs GG	10	F	1.049 (0.956-1.152)	0.314	25.1	0.212	0.504
AA+AG vs GG	10	F	1.043 (0.955-1.140)	0.345	17.5	0.282	0.667
AA vs AG+GG	10	F	0.989 (0.879-1.113)	0.858	0.0	0.797	0.079
<b>Age-unmatched</b>							
A vs G	6	R	0.813 (0.695-0.951)	0.010	66.2	0.011	0.285
AA vs GG	6	R	0.619 (0.434-0.883)	0.008	60.8	0.026	0.405
AG vs GG	6	F	0.812 (0.715-0.921)	0.001	36.3	0.165	0.503
AA+AG vs GG	6	R	0.771 (0.634-0.939)	0.010	61.3	0.024	0.620
AA vs AG+GG	6	R	0.716 (0.528-0.973)	0.033	55.4	0.047	0.628

R: Random-effects model; F: Fixed-effects mode; P<sub>Z</sub>: P value for Z test; P<sub>Q</sub>: P value for Q test.

AMD. Recently, several studies have evaluated the association between *SERPING1* rs2511989 (G>A) polymorphism and AMD. However, the results were inconsistent. Considering these studies' limitations of small sample size, Meta-analysis is a powerful tool for summarizing the contradicting results from different studies with greater statistical power; we undertook to conduct a Meta-analysis of 16 cohort studies involving 9163 cases and 6813 controls.

Among overall comparisons and subgroup comparisons in our study, we only found a significant positive association between rs2511989 polymorphism and AMD in the age-unmatched subgroup under all genetic models. Moreover, meta-regression analysis showed that age matching status was the main source of heterogeneity; not ethnicity or genotyping. These results led us to attach importance to the influence of age on genetic association. When we reviewed the ages of the patients and control

subjects in the age-unmatched subgroups, we found that the control groups were significantly younger than the patient groups (shown in Table 2; some unavailable data can be estimated from the original articles). Ennis *et al*<sup>[18]</sup>. thought that this difference in age might reduce the power of detecting genetic association, because some of the controls could develop AMD in later life. Therefore, they considered the significant positive association detected by their study as reliable. However, we think that this explanation is flawed for the following reasons. Firstly, if it were the case, the pooled ORs of the age-matched subgroup should show relatively greater genetic association than the age-unmatched subgroup. In fact, our study obtained the opposite result. Secondly, the control group population might develop AMD, or die, or not return to follow-up in later life, but we are not sure what will happen in the future. Besides, the age of the control groups in the age-unmatched subgroup might be less than the age of the control groups in the age-matched



**Figure 2** Forest plots for association between the *SERPING1* rs2511989 polymorphism and AMD in different genetic models. The size of the square indicates the relative weight of each study. Bars: 95% confidence interval (95%CI). A: Allele model (A vs G); B: Homozygote model (AA vs GG); C: Heterozygote model (AG vs GG); D: Dominant model (AA+AG vs GG); E: Recessive model (AA vs AG+GG).

subgroup. Taking this into consideration, we do not know whether the power of detecting genetic association is underestimated or overestimated. So the effect of the difference in age on genetic association is complex and indeterminable.

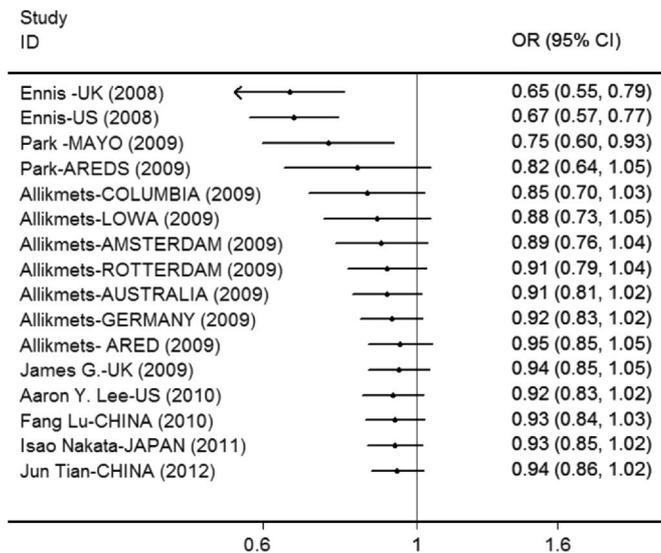
In fact, many studies have proved that genetic association changes with age [31-35]. In particular, the research of Adams *et al* [34], who studied the associations between four variants in the complement factor H and AMD in different age

groups, discovered a change of relationship from protective to risk with increasing age. This finding challenges the conclusion of a strong association between complement factor H and AMD, which has been extensively studied and proved [36,37]. Adams *et al* [34] concluded from their study that it is imperative to ensure that cases and controls have the same age distribution when conducting case-control studies of Adams. Therefore, if the design of the age-unmatched studies is not rigorous, it will lead to a bias of genetic association.

**Table 4 Univariate and multivariate meta-regression analyses of potential source of heterogeneity**

Factors	Coefficient	Standard error	t	P	Adjusted
Ethnicity					
Univariate	0.091	0.130	-0.70	0.495	0.183
Multivariate	-0.288	0.154	-1.88	0.085	
Age matching status					
Univariate	0.235	0.074	-3.19	0.006	0.005
Multivariate	0.285	0.071	-4.02	0.002	
Genotyping					
Univariate	0.000	0.027	0.02	0.987	0.904
Multivariate	0.016	0.030	-0.53	0.605	

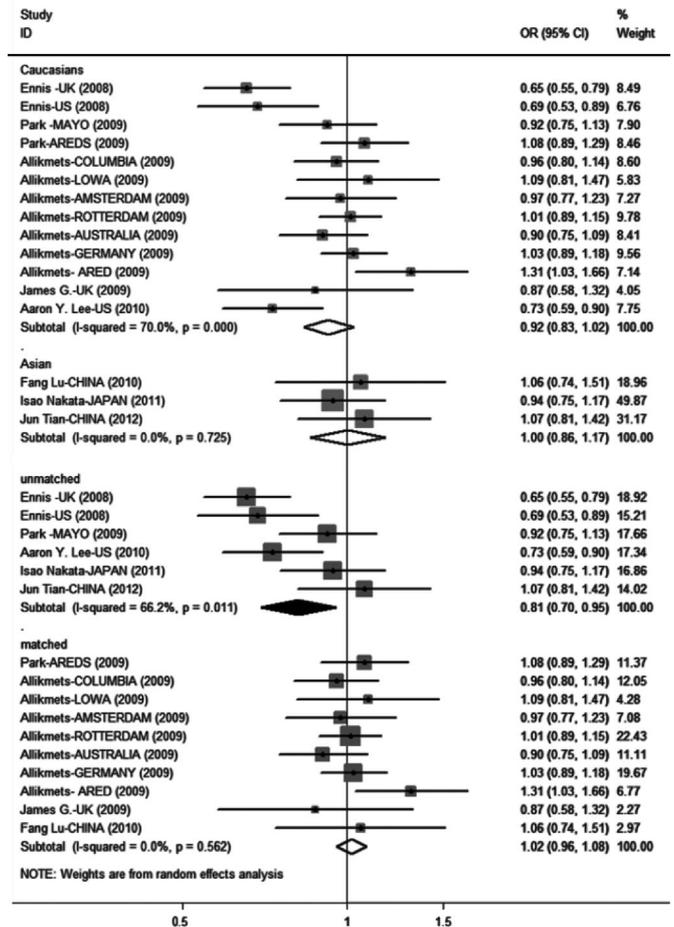
<sup>1</sup>P values were adjusted by Monte Carlo permutation test (10 000 iterations).



**Figure 3 Forest plot of the cumulative Meta-analysis under the allele genetic model** Cumulative odds ratios are shown for each information stepwise accumulated by year.

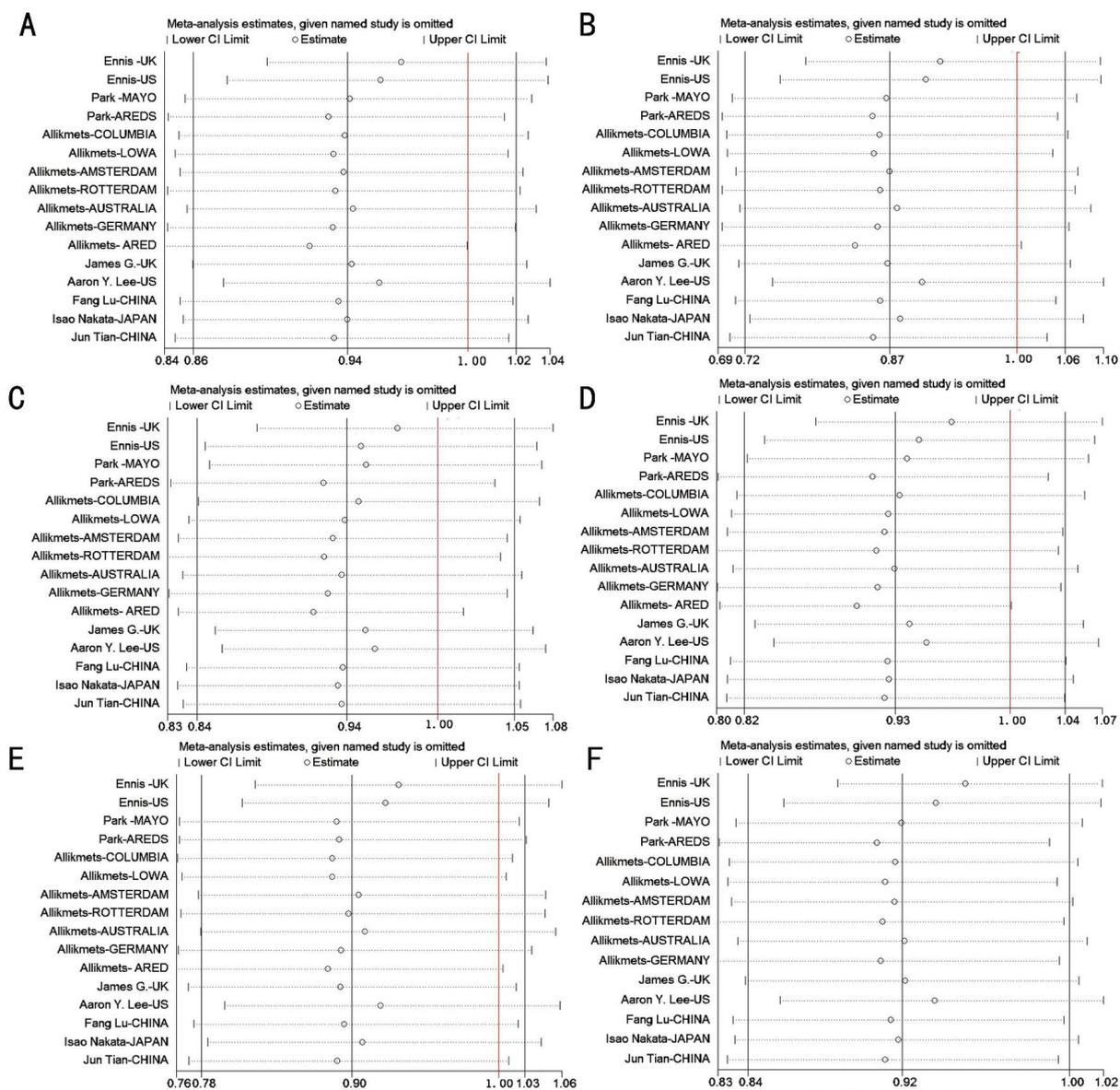
If no account is taken of the influence of age, no association of SNP rs2511989 and AMD was found under overall comparisons, or in other subgroup analyses. Our cumulative Meta-analysis also showed a trend of no association between them as information accumulated year by year. However, after excluding the study that deviated from HWE, significant associations were observed for overall comparisons and white subgroup comparisons under the allele model. This association was not robust when further sensitivity analysis was conducted and did not survive after Bonferroni correction. Hence, we cannot draw any definite conclusion from these analyses.

Although a comprehensive analysis was performed, which included overall and subgroup Meta-analysis, cumulative Meta-analysis, meta-regression and sensitivity analysis, we failed to obtain a determinate conclusion of association between SNP rs2511989 and AMD. This may be due to some inevitable limitations of our study. Firstly, 6 of 16 included studies were age-unmatched. The influence of the age-unmatched studies is complex and indeterminable and



**Figure 4 Subgroup analyses by ethnicity and age matching status for associations between the SERPINC1 rs2511989 polymorphism and AMD under the allele contrast.**

we do not know whether the power of detecting genetic association is underestimated or overestimated. This is a large confounding factor and must be considered in planning further studies. Secondly, this Meta-analysis was limited by the number of cases and controls, especially in the Asian subgroup analysis, which included only three studies. Moreover, some instability results also indicated the number of studies is insufficient. Thus, additional studies are needed to evaluate the relationship of SNP rs2511989 with AMD. Thirdly, owing to a lack of detailed information, such as sex



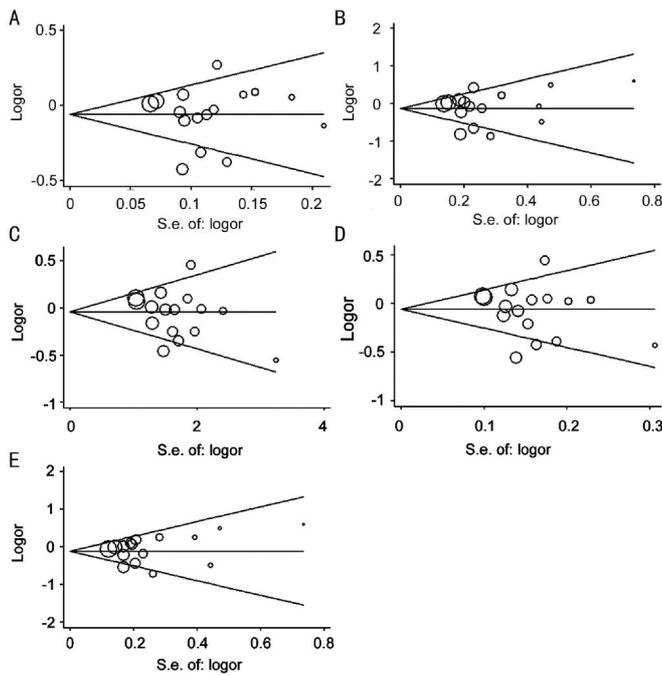
**Figure 5 Sensitivity analysis of the summary odds ratio (OR) coefficients on the relationships between the *SERPINC1* rs2511989 polymorphism and AMD under five genetic models** Results were computed by omitting each study in turn under random-effects model. The red line meant the value of OR is 1.00. The two ends of the dotted lines represented the 95%CI. A: A vs G; B: AA vs GG; C: AG vs GG; D: AA+AG vs GG; E: AA vs AG+GG; F: The result of further sensitivity analysis under allele contrast after exclude the study which deviated from HWE.

and smoking status in individual studies, we failed to perform further subgroup analysis to adjust for these possible confounders. Fourthly, there was significant heterogeneity among included studies, which is mainly caused by age matching status. Although we used the random-effects model to calculate pool ORs, the precision of the outcome would be affected.

In view of the large confounding factor of age-unmatched data and unstable results, we are cautious and do not wish to make a final conclusion on the associations between SNP rs2511989 and AMD. However, judging from all of the analyses and the result of the cumulative Meta-analysis, it would appear that SNP rs2511989 tends to have no association with AMD. The reasons are as follows: 1) the overall and ethnicity subgroup Meta-analysis showed no

relationship between SNP rs2511989 and AMD; 2) the result of the age-matched subgroup analysis, which showed no association between SNP rs2511989 and AMD, is more valid<sup>[38]</sup>; 3) the cumulative Meta-analysis also showed a trend of no association between SNP rs2511989 and AMD as information accumulated by year; 4) after excluding the AREDS cohorts of Allikmets *et al*'s<sup>[20]</sup> study, which deviated from HWE. In controls, we only found an instability and inconspicuous genetic association in one of the five genetic models, which did not survive after Bonferroni correction; 5) there was almost no difference in the mean MAF between cases and controls within all studies, especially in the Asian subgroup.

In summary, our Meta-analysis indicated a tendency of no association between the SNP rs2511989 and AMD. This



**Figure 6** Begg's funnel plot for publication bias test under five genetic models A: A vs G; B: AA vs GG; C: AG vs GG; D: AA+AG vs GG; E: AA vs AG+GG; Each circle represented a separate study for the indicated association, and its size was proportional to the sample size of each study. The shapes of the funnel plots did not reveal any evidence of obvious asymmetry.

conclusion is also supported by two genome-wide association studies, which did not identify any SNP of the *SERPING1* gene associated with AMD<sup>[39,40]</sup>. Be that as it may, the possibility of association cannot be completely ruled out, given current limitations of technology and methodology. The *SERPING1* protein is expressed at the macula, as demonstrated in *in vitro* studies of human donor eyes<sup>[18,41,42]</sup>. Therefore, this gene is worthy of further research in studies with subtler designs, especially where the age of patients and controls is rigorously matched.

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